

ISSN Print: 2394-7500
ISSN Online: 2394-5869
Impact Factor: 8.4
IJAR 2023; 9(2): 323-334
www.allresearchjournal.com
Received: 23-11-2022
Accepted: 16-01-2023

## Joseph Justin Rebello

 Department of Statistics, AC, Mahatma Gandhi University, Kottayam, Kerala, India
## Shafna Noushad

Department of Statistics, AC, Mahatma Gandhi University, Kottayam, Kerala, India

## Surya K S

Department of Statistics, AC, Mahatma Gandhi University, Kottayam, Kerala, India

## Corresponding Author:

Joseph Justin Rebello
Department of Statistics, AC,
Mahatma Gandhi University, Kottayam, Kerala, India

# Some studies on principal component analysis, factor analysis and cluster analysis on a clubbed data 

Joseph Justin Rebello, Shafna Noushad and Surya K S

DOI: https://doi.org/10.22271/allresearch.2023.v9.i2e.10635


#### Abstract

The objective of the study is to apply data reduction techniques PCA and FA on the Healthy Lifestyle Cities Report 2021 and also to cluster the same data using cluster Analysis. The data analysed 44 cities across the globe to uncover where it is easier to lead a well-rounded, healthy lifestyle. From obesity levels to pollution rates, each city has been scored across 6 healthy living metrics. Each of these metrics were awarded a weighted score and these were combined to give each city a total sore out of 100 . This score was then used to rank the 33 cities to determine which were best for healthy living. For the analysis of the data, statistical packages "SPSS" and "R" are being applied.


Keywords: Principle component analysis, factor analysis, cluster analysis, complete linkage method, average linkage method, k means

## 1. Introduction

Multivariate analysis is a statistical procedure for the analysis of data involving more than one type of measurement or observation ${ }^{[1]}$. It may also mean solving problems where more than one dependent variable is analysed simultaneously with other variables. Multivariate analysis is based on the statistical principle of multivariate statistics, which involves observation and analysis of more than one statistical outcome variable at a time. The main advantage of multivariate analysis is that since it considers more than one factor of independent variables that influence the variability of dependent variables, the conclusion drawn is more accurate. The conclusions are more realistic and nearer to the real-life situation ${ }^{[11]}$.

### 1.1 Principal Component Technique

A Principal component analysis is concerned with explaining the variance-covariance structure of a set of variables through a few linear combinations of these variables ${ }^{[8]}$. Principle component may be useful to transform the original set of variables to a new set of uncorrelated variables ${ }^{[2]}$. These new variables are called Principal components which are the normalized linear combination of the original variables and are derived in the decreasing order of importance. PCA is mostly used as a tool in exploratory data analysis and for making predictive models ${ }^{[9]}$. It's often used to visualize genetic distance and relatedness between populations. PCA can be done by eigenvalue decomposition of a data covariance (or correlation) matrix or singular value decomposition of a data matrix, usually after mean centering (and normalizing or using Z scores) the data matrix for each attribute ${ }^{[3]}$. The results of a PCA are usually discussed in terms of component scores, sometimes called factor scores (the transformed variable values corresponding to a particular data point), and loadings (the weight by which each standardized original variable should be multiplied to get the component score).


Fig 1: PCA
PCA is closely related to factor analysis. Factor analysis typically incorporates more domain specific assumptions about the underlying structure and solves eigenvectors of a slightly different matrix. PCA is also related to canonical correlation analysis (CCA)(6). CCA defines coordinate systems that optimally describe the cross-covariance between two datasets while PCA defines a new orthogonal coordinate system that optimally describes variance in a single dataset.

### 1.1.1 Scree Plot

Decision regarding the number of principle components to be taken in any data analysis is decided graphically by a scree plot. The term 'scree' is taken from the word for the rubble at the bottom of the mountain (7).


Fig 2: Scree Plot

### 1.2 Factor Analysis

Factor analysis was developed originally for the analysis of scores on mental tests; however, the methods are useful in a much wider range of situations such as analysing sets of tests of attitudes, sets of physical measurements and sets of economic quantities(4). Factor can be considered as an extension of Principal component analysis. Both can be viewed as attempts to approximate the covariance matrix $\sum$.

### 1.3 Cluster Analysis

Cluster analysis is multivariate method which aims to classify a set of objects in such a way that objects in the same group (called cluster) are more similar to each than to those in other groups (12). Grouping is done on the basis of similarities or distances (dissimilarities) (13).

## 2. Methodology

The variables used for the analysis are the following:

- $X_{1}=$ Rank
- $X_{2}=$ Sunshine Hours(City)
- $X_{3}=$ Life Expectancy (Country)
- $X_{4}=$ Happiness Levels (Country)
- $X_{5}=$ Outdoor Activities (Cities)
- $X_{6}=$ Number of takeout places (City)


### 2.1 Principal Component Analysis

Let $X$ be a $P$ component random variable whose mean is assumed to be $\mu$ and dispersion matrix $\Sigma$, where $\Sigma$ is a real positive matrix. The equation for the characteristic root and the corresponding characteristic vector is given by
$\Sigma \mathrm{X}=\Lambda \mathrm{X}(1)$
According to Hotelling's iterative procedure we start with an initial $P \times 1$ vector $X_{0}$ which is not orthogonal to $e_{1}$, the characteristic vector corresponding to the largest characteristic root $\lambda_{1}$ of $\Sigma$.
Define $X_{i}=\Sigma Z_{i-1} ; i=1,2,3, \ldots \ldots \ldots, P$
$Z_{i}=\frac{\mathrm{xi}}{\sqrt{X i x i}} ; i=1,2,3, \ldots \ldots \ldots, P$
It can be shown that,

$$
\begin{equation*}
\operatorname{Lim}_{i \rightarrow \infty} z_{\mathrm{i}}= \pm e_{i}, \lim _{i \rightarrow \infty} X_{i}^{\prime} X_{i}=\lambda_{1}^{2} \tag{1.2}
\end{equation*}
$$

To find the second characteristic root and the corresponding characteristic vector we define, $\Sigma_{2}=\Sigma-\lambda_{1} e_{1} e_{1}^{\prime}$
Now to find $\lambda_{1}$ and $e_{2}$ we use the same iterative procedure to $\Sigma_{2}$. Repeat the steps (1.1) and (1.2)
Thus the non-zero Eigen values of $\Sigma$ are $\lambda_{1}, \lambda_{2}, \lambda_{3}, \ldots \ldots \lambda_{P}$ with Eigen vectors $e_{1}, e_{2}, e_{3}, \ldots \ldots e_{P}$.
So we get the linear combination $Z_{i}=e_{i}^{\prime} X$
That is,
$Z_{1}=e_{11} X_{1}+e_{12} X_{2} \ldots \ldots+e_{1 P} X_{P}$
$Z_{2}=e_{21} X_{1+} e_{22} X_{2} \ldots \ldots \ldots+e_{2 P} X_{P}$
$Z_{\mathrm{P}}=e_{\mathrm{P} 1} X_{1+} e_{\mathrm{P} 2} X_{2} \ldots \ldots \ldots+e_{\mathrm{P} P} X_{P}$
With the condition $V\left(Z_{1}\right) \geq V\left(Z_{2}\right) \ldots \ldots . V\left(Z_{P}\right) \geq 0$ (1.4)
The linear combination (1.3) is called principal component satisfying (1.4)

### 2.2 Factor Analysis

### 2.2.1 The Orthogonal Factor Model

The observable random vector $X$ with $p$ components has mean $\mu$ and covariance matrix $\sum$. The factor model postulates that $X$ is linearly dependent upon a few unobservable random variables $F_{1}, F_{2}, \ldots, F_{m}$ called common factors and $p$ additional sources of variation $\varepsilon_{1}, \varepsilon_{2}, \varepsilon_{3}, \ldots, \varepsilon_{p}$ called errors or specific factors. In matrix notation,
$X-\mu=L F+\varepsilon$
where $\mathrm{X}-\mu$ is a $\mathrm{p} \times 1$ vector, L is a $\mathrm{p} \times \mathrm{m}$ matrix, F is a mx 1 vector and $\varepsilon$ is a $p \times 1$ vector.
The coefficient $l_{i j}$ is called the loading of the $\mathrm{i}^{\text {th }}$ variable on the $\mathrm{j}^{\text {th }}$ factor, so that the matrix L is the matrix of factor
loadings. The $\mathrm{i}^{\text {th }}$ specific factor $\varepsilon_{i}$ is associated only with the $i^{\text {th }}$ response $X_{i}$. The $p$ deviations $X_{1}-\mu_{1}, X_{2}-\mu_{2}$, $\ldots X_{p}-\mu_{p}$, are expressed in terms of $p+m$ random variables $F_{1}, F_{2}, \ldots . F_{m}, \varepsilon_{1}, \varepsilon_{2} \ldots . \varepsilon_{p}$ Which are unobservable. This distinguishes the factor model expressed in equation (2) from the regression model in the independent variables observed.
With so many unobservable quantities, a direct verification of the factor model from observations on $X_{1}, X_{2}, \ldots . X_{P}$ is hope less. However, with some additional assumptions about the random vectors $F$ and $\varepsilon$, the mode
$X-\mu=L F+\varepsilon$,
Implies certain covariance relationships can be checked.
The assumptions are:
$E(F)=0_{m \times 1}, \operatorname{Cov}(F)=E\left(F F^{\prime}\right)=I_{m \times m}$
$E(\varepsilon)=0_{p \times 1}, \operatorname{Cov}(\varepsilon)=\varphi_{p \times p}=\left[\begin{array}{ccc}\varphi_{1} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \varphi_{p}\end{array}\right]$

## $F$ and $\varepsilon$ are independent.

$i e, \operatorname{Cov}(F, \varepsilon)=E\left(\varepsilon F^{\prime}\right)=0_{p \times m}$
These assumptions and the relation $X-\mu=L F+\varepsilon$ constitute the orthogonal factor model. The factor analysis model with the above assumptions is called orthogonal Factor analysis model.

### 2.2.2 Methods of Estimation

Given observations $X_{1}, X_{2}, \ldots X_{n}$ on $p$ generally correlated variables. The sample covariance matrix $S$ is an estimator of the unknown population covariance matrix $\sum$. If the offdiagonal elements of S are small or those of the sample correlation matrix R essentially zero, the variables are not related, and a factor analysis will not be useful. In such circumstances, the specific factors play a dominant role, but the aim of factor analysis is to determine a few important common factors. If $\sum$ appears to deviate significantly from a diagonal matrix, then a factor model can be entertained and the initial problem is one of estimating the factor loadings $I_{i j}$ and specific variance $\varphi_{i}$.

### 2.2.3 Factor Rotations

The results of factor extraction, unaccompanied by rotation are likely to be hard to interpret regardless, of which method of extraction is used. After extraction, rotation is used to improve the interpretability and scientific utility of solution. It is not used to improve the quality of mathematical fit between observed and reproduced correlation matrices because all orthogonally rotated solutions are equivalent to one another and to the solution before rotation. All factor loadings obtained from the initial loadings by an orthogonal transformation have the same ability to reproduce the covariance (or correlation) matrix. We know that, an orthogonal transformation corresponds to a rigid rotation of the co-ordinate axes. For this reason, an orthogonal transformation of the factor loadings, as well as the implied orthogonal transformations of the factors, is called factor rotation. Rotations are ordinarily used after extraction to maximize high correlations and minimize low ones.

If $\hat{L}$ is the $p \times m$ matrix of estimated factor loadings obtained by any method (principal component or maximum likelihood) then

$$
L^{\wedge *}=L^{\wedge} T, \text { where } T T^{\prime}=T^{\prime} T=I \text { (orthogonal). }
$$

Hence $L^{\wedge *}=L^{\wedge} T$ is a $p \times m$ matrix of "rotated" loadings. Moreover, the estimated covariance (or correlation) matrix remains unchanged since

$$
L^{\wedge} L^{\wedge \prime}+\Psi=L^{\wedge} T T^{\prime} L^{\wedge \prime}+\Psi^{\wedge}=L^{\wedge *} L^{\wedge * \prime}+\Psi^{\wedge} .
$$

Hence the residual matrix
$S_{n}-L^{\wedge} L^{\wedge}-\Psi^{\wedge}=S_{n}-L^{\wedge *} L^{\wedge}{ }^{* \prime}+\Psi^{\wedge}$
remains unchanged. Moreover, the specific variances $\Psi_{i}^{\hat{a}}$ and hence the communalities $h_{i}^{\wedge}$ are unaltered.

### 2.2.4 Factor Scores

Usually in Factor analysis, the interest in centered on the parameters in the factor model. We may also require the estimated values of the common factors called factor scores. These quantities are often used for diagnostic purposes, as well as inputs to a subsequent analysis. Factor scores are not estimates of unknown parameters in the usual sense. Rather, they are estimates of the values for the unobserved random vectors
$F_{j}, j=1,2, \ldots, n$. That is, factor scores
$f_{j}^{\wedge}=$ Estimates of the values $f_{j}$ attained by $F_{j}\left(j^{t h}\right.$ case $)$.
The estimation of factor scores is done using weighted least squares method as follows:
Suppose the mean vector $\mu$, the factor loading $L$ and specific variance $\Psi$ are known for the factor model, $X-\mu=$ $L F+\varepsilon$.
Further, regard the specific factor $\varepsilon^{\prime}=\left(\varepsilon_{1}, \varepsilon_{2}, \ldots \ldots, \varepsilon_{p}\right)$ as errors. Since
$V\left(\varepsilon_{i}\right)=\Psi_{i}, i=1,2, \ldots p$ need not be equal. Bartlett suggested that weighted least squares can be used to estimate the common factor values.
The sum of squares of the errors weighted, by the reciprocal of their variance is
$\sum_{i=1}^{p} \frac{\varepsilon_{i}^{2}}{\Psi_{i}}=\varepsilon^{\prime \Psi^{-1}} \varepsilon=(X-L F-\mu)^{\prime} \Psi^{-1}(X-L F-\mu)$.
If we take $L^{\wedge}, \Psi^{\wedge}$ and $\mu^{\wedge}=\bar{X}$, the estimates of $L, \Psi$ and $\mu$ as the true values, then the factor scores for the $j^{t h}$ case is obtained by minimizing

$$
\left(X_{j}-L F_{j}-\hat{\mu}\right)^{\prime} \Psi^{-1}\left(X_{j}-L F_{j}-\hat{\mu}\right)
$$

The solution is given by

$$
F_{j}^{\wedge}=\left(L^{\wedge \prime} \Psi^{-1} L^{\wedge}\right)^{-1} L^{\wedge \prime} \Psi^{\wedge-1}\left(X_{j}-\bar{X}\right), \mathrm{j}=1,2 \ldots \mathrm{n}
$$

The factor scores generated have sample mean vector $\mathbf{o}$ and zero sample covariance matrix.

## 3. Analysis of Data

### 3.1 Principal component analysis

The result obtained by using the principal component analysis as the extraction method is given below.
Table 1: Communalities

| Communalities |  |  |
| :---: | :---: | :---: |
|  | Initial | Extraction |
| Rank | 1.000 | .963 |
| Sunshine Hours | 1.000 | .997 |
| Life expectancy | 1.000 | .952 |
| Happiness levels | 1.000 | .985 |
| Outdoor activities | 1.000 | .963 |
| Number of take out places | 1.000 | .931 |

## Extraction

It indicates that proportion of variance that can be explained by the principal components. Now, a scree plot displays the eigen values associated with a component or factor in
descending order versus the number of the components or factor. We use scree plots in principal components analysis and the factor analysis to visually assess which components or factors explain most of the variability in the data.


Fig 3: Scree plot
From the scree plot, it can be concluded that we can extract 5 principal components
Table 2: Total Variance Explained

| Total Variance Explained |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  |
|  | Total | \% of Variance | Cumulative \% | Total | \% of Variance | Cumulative \% |
| 1 | 2.078 | 34.625 | 34.625 | 2.078 | 34.625 | 34.625 |
| 2 | 1.624 | 27.071 | 61.697 | 1.624 | 27.071 | 61.697 |
| 3 | 1.119 | 18.654 | 80.350 | 1.119 | 18.654 | 80.350 |
| 4 | .688 | 11.464 | 91.814 | .688 | 11.464 | 91.814 |
| 5 | .282 | 4.699 | 96.514 | .282 | 4.699 | 96.514 |
| 6 | .209 | 3.486 | 100.000 |  |  |  |

The principal components with eigen values greater than 1 are normally considered. Here 5 components have eigen value greater than one. So these components are considered.

Also from the scree plot, 5 components are retained. Therefore 5 components are extracted (14).

Table 3: Component Matrix

| Component Matrix |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Component |  |  |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ |
| Rank | -.437 | .401 | -.699 | .338 | .089 |
| Sunshine Hours | -.547 | -.286 | .455 | .636 | -.075 |
| Life expectancy | .895 | .108 | .086 | .213 | -.292 |
| Happiness levels | .872 | .014 | -.036 | .331 | .337 |
| Outdoor activities | -.158 | .692 | .643 | -.086 | .196 |
| Number of take out places | .021 | .944 | -.046 | -0.83 | -.176 |

It is seen that the $1^{\text {st }}$ component explains $34.625 \%$ of variation of the data set, $2^{\text {nd }}$ component explains $27.071 \%$ of data set, $3^{\text {rd }}$ component explains $18.654 \%$ of variation of the data set, $4^{\text {th }}$ component explains $11.464 \%$ of variation of the data set and $5^{\text {th }}$ component explains $4.699 \%$ of data set. That is, first five component explain $96.514 \%$ of variation of data set. It is also from the component matrix that the first 5 components are highly influenced by all the 6 factors.

Also we get, the first component is highly influenced by the variables $X_{3}$ followed by $X_{4}$ and $X_{6}$, second component is highly influenced by the variables $X_{6}$ followed by $X_{5}$ and $X_{1}, 3^{\text {rd }}$ component is highly influenced by the variables $X_{5}$ followed by $X_{2}$ and $X_{3}, 4^{\text {th }}$ component is highly influenced by the variables $X_{2}$ followed by $X_{1}$ and $X_{4}$ and the $5^{\text {th }}$ component is highly influenced by the variables $X_{4}$ followed by $X_{5}$ and $X_{1}$.
From the component matrix,

Table 4: List of variables contributing more towards variability

| Components |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ | $\mathbf{5}$ |  |
| $\mathrm{X}_{3}$ (Life expectancy) | $\mathrm{X}_{6}$ (No: of take out places) | $\mathrm{X}_{5}$ (Outdoor Activities) | $\mathrm{X}_{2}$ (Sunshine hours) | $\mathrm{X}_{4}$ (Happiness levels) |  |
| $\mathrm{X}_{4}$ (Happiness levels) | $\mathrm{X}_{5}$ (Outdoor Activities) | $\mathrm{X}_{2}$ (Sunshine hours) | $\mathrm{X}_{1}$ (Rank) | $\mathrm{X}_{5}$ (Outdoor Activities) |  |
| $\mathrm{X}_{6}$ (No: of take out places) | $\mathrm{X}_{1}$ (Rank) | $\mathrm{X}_{3}$ (Life expectancy) | $\mathrm{X}_{4}$ (Happiness levels) | $\mathrm{X}_{1}$ (Rank) |  |

Let the principal components be $U_{1}, U_{2}, U_{3}, U_{4}$ and $U_{5}$
From the score coefficient matrix we get,
$U_{1}=-.437 X_{1}-.547 X_{2}+.895 X_{3}+.872 X_{4}-.158 X_{5}+.021 X_{6}$
$U_{2}=.401 X_{1}-.286 X_{2}+.108 X_{3}+.014 X_{4}+.692 X_{5}+.944 X_{6}$
$U_{3}=-.699 X_{1}+.455 X_{2}+.086 X_{3}-.036 X_{4}+.643 X_{5}-.046 X_{6}$
$U_{4}=.338 X_{1}+.636 X_{2}+.213 X_{3}+331 X_{4}-.086 X_{5}+.083 X_{6}$
$U_{5}=.089 X_{1}-.075 X_{2}-292 X_{3}+.337 X_{4}+.196 X_{5}-.176 X_{6}$

### 3.2 Factor analysis

Table 5: Communalities

| Communalities |  |  |
| :---: | :---: | :---: |
|  | Initial | Extraction |
| Rank | 1.000 | .729 |
| Sunshine hours | 1.000 | .999 |
| Life expectancy | 1.000 | .869 |
| Happiness levels | 1.000 | .816 |
| Number of take out places | 1.000 | .787 |

One or more communality estimates greater than 1 were encountered during iterations. The resulting solution should be interpreted with caution.

Table 6: Total Variance Explained

| Total Variance Explained |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Component | Initial Eigenvalues |  |  | Extraction Sums of Squared Loadings |  |  | Rotation Sums of Squared Loadings |  |  |
|  | Total | \% of Variance | Cum \% | Total | \% of Variance | Cum \% | Total | \% of Variance | Cum \% |
| 1 | 2.068 | 41.355 | 41.355 | 2.068 | 41.355 | 41.355 | 1.827 | 36.541 | 36.541 |
| 2 | 1.424 | 28.481 | 69.836 | 1.424 | 28.481 | 69.836 | 1.357 | 27.136 | 63.676 |
| 3 | . 707 | 14.149 | 83.985 | . 707 | 14.149 | 83.985 | 1.015 | 20.309 | 83.985 |
| 4 | . 553 | 11.064 | 95.049 |  |  |  |  |  |  |
| 5 | . 248 | 4.951 | 100.000 |  |  |  |  |  |  |



Fig 4: Scree plot
Table 7: Component Matrix

| Component Matrix |  |  |  |
| :---: | :---: | :---: | :---: |
|  | Component |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ |
| Rank | -.425 | .728 | .133 |
| Sunshine hours | -.555 | -.423 | .715 |
| Life expectancy | .903 | .004 | .231 |
| Happiness levels | .867 | -.022 | .252 |
| Number of take-out places | .106 | .845 | .249 |
| Extraction Method: Principal Component Analysis. |  |  |  |
| a. 3 components extracted. |  |  |  |

Table 8: Rotated Component Matrix

| Rotated Component Matrix |  |  |  |
| :---: | :---: | :---: | :---: |
|  | Component |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ |
| Rank | -.346 | .780 | .025 |
| Sunshine hours | -.194 | -.088 | .976 |
| Life expectancy | .916 | -.038 | -.168 |
| Happiness levels | .893 | -.050 | -.126 |
| Number of take-out places | .182 | .858 | -.131 |

Extraction Method: Principal Component Analysis. Rotation Method: Varimax with Kaiser Normalization. a. Rotation converged in 4 iterations.

SPSS calculates the factor loadings for each variable in the analysis. The loadings of a factor explain each variable. Large loadings (positive or negative) indicate the high influences the variable. Small loadings (positive or negative) indicate that the influence on the variable (5).
Unrotated factor loadings are often difficult to interpret. Factor rotations simplify structure, and make the factor loadings easier to interpret.

Table 9: Component Transformation Matrix

| Component Transformation Matrix |  |  |  |
| :---: | :---: | :---: | :---: |
| Component | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ |
| 1 | .907 | -.130 | -.400 |
| 2 | -.022 | .935 | -.354 |
| 3 | .420 | .330 | .845 |

Extraction Method: Principal Component Analysis. Rotation Method: Varimax with Kaiser Normalization.

Interpret the output in the same way as PCA although the Component matrix is called the Factor Matrix if the extraction method has changed.
Look at the Rotated Factor matrix to see which variables contribute most to each factor (PC). Variables measuring the same underlying latent variable should all have high loadings on a particular factor and by looking at the raw variables, a sensible name can be given to the factor. The next factor should be measuring another latent variables etc. The factor plot is useful for assessing grouping of variables on more than one factor. If there are two factors, the variables appear on a scatterplot.
Using the scree plot, it can be concluded that we can extract 3 Factors.
The Factors are
$Y_{1}=-.346 X_{1}-.194 X_{2}+.916 X_{3}+.893 X_{4}+.182 X_{5}$
$Y_{2}=.780 X_{1}-.088 X_{2}-.038 X_{3}-.050 X_{4}+.858 X_{5}$
$Y_{3}=.025 X_{1}+.976 X_{2}-.168 X_{3}-.126 X_{4}-.131 X_{5}$
It is seen that the $1^{\text {st }}$ component explains $36.541 \%$ of variation of the data set, $2^{\text {nd }}$ component explains $27.136 \%$ of data set and $3^{\text {rd }}$ component explains $20.309 \%$ of variation of the data set. That is, first 3 component explain $83.985 \%$ of variation of data set. It is also from the component matrix that the first 3 components are highly influenced by all the 5 factors.

### 3.3 Cluster analysis

Here we are comparing average linkage cluster and complete linkage cluster

### 3.3.1 Complete Linkage Method

Table 10: Cluster Membership

| Cluster 1 | Cluster 2 | Cluster 3 |
| :---: | :---: | :---: |
| Amsterdam | Vienna | Jakarta |
| Sydney | Stockholm | Cairo |
| Barcelona | Copenhagen | Mumbai |
| Tokyo | Helsinki | Johannesburg |
| Paris | Fukuoka |  |
| London | Berlin |  |
| New York | Vancouver |  |
|  | Melbourne |  |
|  | Beijing |  |
|  | Bangkok |  |
|  | Buenos Aires |  |
|  | Toronto |  |
|  | Madrid |  |
|  | Seoul |  |
|  | Frankfurt |  |
|  | Geneva |  |
|  | Tel Aviv |  |
|  | Istanbul |  |
|  | Taipei |  |
|  | Los Angeles |  |
|  | Boston |  |
|  | Dublin |  |
|  | Chicago |  |
|  | Hong Kong |  |
|  | Shanghai |  |
|  | Brussels |  |
|  | San Francisco |  |
|  | Sao Paulo |  |
|  | Zurich |  |
|  | Milan |  |
|  | Washington, D.C |  |
|  | Moscow |  |
|  | Mexico City |  |

The cluster plot is given in Figure 3


Fig 5: Cluster plot

### 3.3.2 Average Linkage method

Table 11: Cluster Membership

| Cluster 1 | Cluster 2 | Cluster 3 |
| :---: | :---: | :---: |
| Amsterdam | Vienna | Jakarta |
| Sydney | Stockholm | Cairo |
| Barcelona | Copenhagen | Mumbai |
|  | Helsinki | Johannesburg |
|  | Fukuoka |  |
|  | Berlin |  |
|  | Vancouver |  |
|  | Melbourne |  |
|  | Beijing |  |
|  | Bangkok |  |
|  | Buenos Aires |  |
|  | Toronto |  |
|  | Madrid |  |
|  | Seoul |  |
|  | Frankfurt |  |
|  | Geneva |  |
|  | Tel Aviv |  |
|  | Istanbul |  |
|  | Taipei |  |
|  | Los Angeles |  |
|  | Boston |  |
|  | Dublin |  |
|  | Tokyo |  |
|  | Chicago |  |
|  | Hong Kong |  |
|  | Shanghai |  |
|  | Brussels |  |
|  | San Francisco |  |
|  | Paris |  |
|  | Sao Paulo |  |
|  | Zurich |  |
|  | London |  |
|  | Milan |  |
|  | Washington, D.C. |  |
|  | New York |  |
|  | Moscow |  |
|  | Mexico City |  |

The cluster plot is given in Figure 4


Fig 6: Cluster plot

$$
\text { ~ } 330 \text { ~ }
$$

Table 12: Final Clusters

| Final clusters |  |
| :---: | :---: |
|  | Member Average Linkage |
| Member Complete Linkage | 123 |
|  | 1340 |
|  | 20330 |
|  | 3004 |

This table tell us that using average linkage method, there are 3 observations belong to cluster 1. Four plus 33 observations belong to cluster 2 and 4 observations belong to cluster 3.
Using complete linkage method, there are 3 plus 4 Cities that belong to cluster 1,33 belong to cluster 2 and 4 belong to cluster 3.

If we compare Average linkage method and Average linkage there are good match for 3 cities, both methods listed them as cluster one. Whereas, there are 33 cities both indicated belong to cluster 2 and also 4 cities belong to cluster 3.
But also we can see there is some mismatch, 4 cities have membership in cluster 2 based on Average method. But this cities have membership in cluster 1 if you use complete linkage method.
So this table allow us to compare these two different methods, Complete linkage and Average Linkage Method. We can also calculate Cluster means,

## Average Value of Complete linkage

Table 13: Average value of Complete Linkage

| Group | Rank | Sunshine hours | Life Expectancy | Happiness levels | Outdoor Activities |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | -1.44022305 | 0.1867216 | 0.6896830 | 0.5901925 | 2.02077308 |
| 2 | 0.08521407 | -0.1686275 | 0.1983321 | 0.1615732 | -0.15579106 |
| 3 | 0.29193710 | 1.4197628 | -2.3518346 | -2.0000969 | -0.07451255 |

The average value help us to find out which is important variable. For example in case of outdoor activities 2.02077308 is the highest value which means most of outdoor activities occurs at cluster one and -0.15579106 the lowest value which belongs to cluster 2 indicates minimal
outdoor activities occurs at cluster 2. These Averages indicate which variables are really playing an important role in characterizing the clusters.

## Average Value of Average linkage method

Table 14: Average value of Average Linkage

| Group | Rank | Sunshine hours | Life Expectancy | Happiness levels | Outdoor Activities |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | -0.01668212 | -0.2531972 | 0.6052961 | 0.381211 | 1.60070509 |
| 2 | -0.03184768 | -0.1183840 | 0.1566747 | 0.161573 | -0.33051168 |
| 3 | 0.29193710 | 1.4197628 | -2.3518346 | -2.000097 | -0.07451255 |

The average value help us to find out which is important variable. For example in case of outdoor activities 1.60070509 is the highest value which means most of outdoor activities occurs at cluster one and -0.33051168 the lowest value which belongs to cluster 2 indicates minimal outdoor activities occurs at cluster 2. These Averages
indicate which variables are really playing an important role in characterizing the clusters.

## K-mean Cluster

The important step in K-means clustering technique is to decide the number of clusters.
So here we choose $\mathrm{k}=4$

Table 15: Initial Clusters Centers

| Initial Cluster Centers |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Cluster |  |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| Rank | 38.0000 | 36.0000 | 23.0000 | 37.0000 |
| Sunshine hours | 1633.0000 | 2003.0000 | 3542.0000 | 1566.0000 |
| Life expectancy | 80.40000 | 73.90000 | 70.70000 | 82.60000 |
| Happiness levels | 7.1600 | 6.3700 | 4.1500 | 7.5600 |
| Outdoor activities | 433.0000 | 158.0000 | 323.0000 | 69.0000 |
| Number of take out places | 6417.00000 | 3355.00000 | 250.00000 | 538.00000 |

Table 16: Iteration History

| Iteration History |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Iteration | Change in Cluster Centers |  |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| 1 | 331.650 | 400.172 | 723.467 | 596.835 |
| 2 | .000 | .000 | 78.415 | 51.946 |
| 3 | .000 | .000 | 49.468 | 34.404 |
| 4 | .000 | .000 | 146.533 | 146.671 |
| 5 | .000 | .000 | 26.163 | 33.417 |
| 6 | .000 | .000 | .000 | .000 |

Convergence achieved due to no or small change in cluster centers. The maximum absolute coordinate change for any
center is. 000 . The current iteration is 6 . The minimum distance between initial centers is 2013.054.

Table 17: Cluster Membership

| Cluster Membership |  |  |  |
| :---: | :---: | :---: | :---: |
| Case Number | City | Cluster | Distance |
| 1 | Amsterdam | 4 | 344.920 |
| 2 | Sydney | 3 | 341.566 |
| 3 | Vienna | 4 | 222.179 |
| 4 | Stockholm | 4 | 239.395 |
| 5 | Copenhagen | 4 | 332.622 |
| 6 | Helsinki | 4 | 530.842 |
| 7 | Fukuoka | 3 | 386.907 |
| 8 | Berlin | 4 | 917.351 |
| 9 | Barcelona | 2 | 845.329 |
| 10 | Vancouver | 4 | 181.812 |
| 11 | Melbourne | 3 | 444.772 |
| 12 | Beijing | 3 | 641.573 |
| 13 | Bangkok | 3 | 941.124 |
| 14 | Buenos aires | 3 | 612.666 |
| 15 | Toronto | 4 | 882.962 |
| 16 | Madrid | 2 | 793.143 |
| 17 | Jakarta | 3 | 209.812 |
| 18 | Seoul | 4 | 531.111 |
| 19 | Frankfurt | 4 | 330.863 |
| 20 | Geneva | 3 | 480.639 |
| 21 | Tel aviv | 3 | 697.151 |
| 22 | Istanbul | 4 | 525.190 |
| 23 | Cairo | 3 | 988.630 |
| 24 | Taipei | 4 | 378.890 |
| 25 | Los angeles | 3 | 714.559 |
| 26 | Mumbai | 3 | 364.089 |
| 27 | Boston | 3 | 359.783 |
| 28 | Dublin | 4 | 354.732 |
| 29 | Tokyo | 1 | 331.650 |
| 30 | Chicago | 3 | 520.336 |
| 31 | Hong kong | 4 | 448.820 |
| 32 | Shanghai | 4 | 484.849 |
| 33 | Brussels | 4 | 298.762 |
| 34 | San francisco | 3 | 303.000 |
| 35 | Paris | 2 | 1433.152 |
| 36 | Sao paulo | 2 | 400.172 |
| 37 | Zurich | 4 | 366.522 |
| 38 | London | 1 | 331.650 |
| 39 | Johannesburg | 3 | 514.118 |
| 40 | Milan | 2 | 721.890 |
| 41 | Washington,d.c. | 3 | 327.673 |
| 42 | New york | 2 | 347.401 |
| 43 | Moscow | 2 | 343.325 |
| 44 | Mexico city | 3 | 489.337 |

Table 18: Final Cluster Centers

| Final Cluster Centers |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Cluster |  |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| Rank | 33.5000 | 31.5714 | 22.5556 | 17.4118 |
| Sunshine hours | 1755.0000 | 2196.5714 | 2798.5000 | 1765.4706 |
| Life expectancy | 81.80000 | 78.72857 | 76.13889 | 79.67647 |
| Happiness levels | 6.5150 | 6.3843 | 6.1689 | 6.7282 |
| Outdoor activities | 410.0000 | 297.2857 | 196.1111 | 175.5294 |
| Number of take-out places | 6109.50000 | 3033.71429 | 889.11111 | 825.76471 |

Table 19: Distance between Final Centers

| Distances between Final Cluster Centers |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Cluster | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| 1 |  | 3109.367 | 5327.969 | 5288.970 |
| 2 | 3109.367 |  | 2229.790 | 2252.979 |
| 3 | 5327.969 | 2229.790 |  | 1035.193 |
| 4 | 5288.970 | 2252.979 | 1035.193 |  |

The initial cluster centres are given in the table 20 followed by the changes to clusters centres in the iteration history. The last row should show negligible change, The final cluster centres show how the variables differ in each cluster. It should be clear which variables are most different and therefore define each cluster but ANOVA table shows which variables contribute most to the separation(Highest Fstatistics) and least.
The F tests should be used only for descriptive purposes because the clusters have been chosen to maximize the differences among cases in different clusters. The observed significance levels are not corrected for this and thus cannot be interpreted as tests of the hypothesis that the cluster means are equal.

Table 20: Number of cases in each Cluster

| Number of Cases in each Cluster |  |  |
| :---: | :---: | :---: |
| Cluster | 1 | 2.000 |
|  | 2 | 7.000 |
|  | 3 | 18.000 |
|  | 4 | 17.000 |
| Valid | 44.000 |  |
| Missing |  | 1.000 |

## 4. Summary and Conclusions

### 4.1 Principal component

Using the scree plot, it can be concluded that we can extract three principal components.
The components are,
$U_{1}=-.437 X_{1}-.547 X_{2}+.895 X_{3}+.872 X_{4}-.158 X_{5}+.021 X_{6}$
$U_{2}=.401 X_{1}-.286 X_{2}+.108 X_{3}+.014 X_{4}+.692 X_{5}+.944 X_{6}$
$U_{3}=-.699 X_{1}+.455 X_{2}+.086 X_{3}-.036 X_{4}+.643 X_{5}-.046 X_{6}$
$U_{4}=.338 X_{1}+.636 X_{2}+.213 X_{3}+331 X_{4}-.086 X_{5}+.083 X_{6}$
$U_{5}=.089 X_{1}-.075 X_{2}-292 X_{3}+.337 X_{4}+.196 X_{5}-.176 X_{6}$

- It is seen that the $1^{\text {st }}$ component explains $34.625 \%$ of variation of the data set, $2^{\text {nd }}$ component explains $27.071 \%$ of data set, $3^{\text {rd }}$ component explains $18.654 \%$ of variation of the data set, $4^{\text {th }}$ component explains $11.464 \%$ of variation of the data set and $5^{\text {th }}$ component explains $4.699 \%$ of data set. That is, first five component explain $96.514 \%$ of variation of data set
- Also we get, the first component is highly influenced by the variables $X_{3}$ followed by $X_{4}$ and $X_{6}$, second component is highly influenced by the variables $X_{6}$ followed by $X_{5}$ and $X_{1}, 3^{\text {rd }}$ component is highly influenced by the variables $X_{5}$ followed by $X_{2}$ and $X_{3}, 4^{\text {th }}$ component is highly influenced by the variables $X_{2}$ followed by $X_{1}$ and $X_{4}$ and the $5^{\text {th }}$ component is
highly influenced by the variables $X_{4}$ followed by $X_{5}$ and $X_{1}$.


### 4.2 Factor Analysis

Using the scree plot, it can be concluded that we can extract 3 Factors.
The Factors are,
$Y_{1}=-.346 X_{1}-.194 X_{2}+.916 X_{3}+.893 X_{4}+.182 X_{5}$
$Y_{2}=.780 X_{1}-.088 X_{2}-.038 X_{3}-.050 X_{4}+.858 X_{5}$
$Y_{3}=.025 X_{1}+.976 X_{2}-.168 X_{3}-.126 X_{4}-.131 X_{5}$

- It is seen that the $1^{\text {st }}$ component explains $36.541 \%$ of variation of the data set, $2^{\text {nd }}$ component explains $27.136 \%$ of data set and $3^{\text {rd }}$ component explains $20.309 \%$ of variation of the data set. That is, first 3 component explain $83.985 \%$ of variation of data set. It is also from the component matrix that the first 3 components are highly influenced by all the 5 factors.


### 4.3 Cluster Analysis

Complete Linkage and Average Linkage Method
Table 21: Final Clusters

| Final clusters |  |
| :---: | :---: |
| Member Average Linkage |  |
|  | 123 |
| Member Complete Linkage | 1340 |
|  | 20330 |
|  | 3004 |

This table tell us that using average linkage method, there are 3 observations belong to cluster 1. Four plus 33 observations belong to cluster 2 and 4 observations belong to cluster 3.
Using complete linkage method, there are 3 plus 4 Cities that belong to cluster 1,33 belong to cluster 2 and 4 belong to cluster 3.
If we compare Average linkage method and Average linkage there are good match for 3 cities, both methods listed them as cluster one. Whereas, there are 33 cities both indicated belong to cluster 2 and also 4 cities belong to cluster 3.
But also we can see there is some mismatch, 4 cities have membership in cluster 2 based on Average method. But this cities have membership in cluster 1 if you use complete linkage method.
So this table allow us to compare these two different methods, Complete linkage and Average Linkage Method.
K-means

Table 22: The final clusters are,

| Final Cluster Centers |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Cluster |  |  |  |
|  | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ | $\mathbf{4}$ |
| Rank | 33.5000 | 31.5714 | 22.5556 | 17.4118 |
| Sunshine | 1755.0000 | 2196.5714 | 2798.5000 | 1765.4706 |
| Life expectancy | 81.80000 | 78.72857 | 76.13889 | 79.67647 |
| Happiness levels | 6.5150 | 6.3843 | 6.1689 | 6.7282 |
| Outdoor activities | 410.0000 | 297.2857 | 196.1111 | 175.5294 |
| Number of take out places | 6109.50000 | 3033.71429 | 889.11111 | 825.76471 |

## 5. References

1. Bartlett MS. Multivariate Analysis, Journal of Royal Statistics Society. 1947;2:176-197.
2. Cadima J, Cerdeira JO, Minhoto M. Computational aspects of algorithms for variable selection in the context of principal components. Comp. Stat. Data Anal. 2004;47:225-236.
3. Cadima J, Jolliffe IT. On relationships between uncentred and column-centred principal component analysis. Pak. J Stat. 2009;25:473-503.
4. Gorsuch RL. Factor Analysis (2nd ed). Hillsdae, NJ: Erlbaum; c1983.
5. Hallin M, Paindaveine D, Verdebout T. Efficient Restimation of principal and common principal components. J Am. Stat. Assoc. 2014;109:1071-1083.
6. Hotelling H. Analysis of a complex of statistical variables into principal components. J Educ. Psychol. 1933;24:417-441, 498-520.
7. Jackson JE. A user's guide to principal components. New York, NY: Wiley; c1991.
8. Jolliffe IT. Principal component analysis, 2nd edn. New York, NY: Springer-Verlag; c2002.
9. Jolliffe IT, Cadima J. Principal component analysis: a review and recent developments, Phil. Trans. R. Soc. A. 2016;374:20150202.
10. Li Y, Wang N, Carroll RJ. Selecting the number of principal components in functional data. J Am. Stat. Assoc. 2013;108:1284-1294.
11. Rao CR. Tests of significance in Multivariate Analysis, Biometrika. 1948;35:58-79.
12. Romsberg H. Cluster Analysis for Researchers, Lulu Press; c2004.
13. Sharma M, Wadhawan P. A Cluster Analysis Study of Small and Medium Enterprises, The IUP Journal of Management Research. 2009, 8(10).
14. SPSS Inc. SPSS 16.0 [Computer software]. Chicago: Author; c2007.
